## Marginal false discovery rates

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## Where we're at and where we're going

- At this point, we've covered the most widely used approaches to fitting penalized regression models in the standard setting
- The remainder of the course will focus on:
  - Inference for  $\beta$
  - Other models, such as logistic regression and Cox regression
  - Other covariate structures, such as grouping and fusion
- We'll begin with inference

### Inferential questions

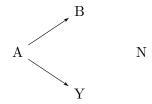
- Up until this point, our inference has been restricted to the predictive ability of the model (which we can obtain via cross-validation)
- This is useful, of course, but we would also like to be able to ask the questions:
  - How reliable are the selections made by the model? What is its false discovery rate?
  - How accurate are the estimates yielded by the model? Can we obtain confidence intervals for  $\beta$ ? Even for  $\beta_j$  not selected by the model?

#### Overview

- As I've remarked previously, little progress was made on these questions until relatively recently, and the field is still very much unsettled as far as a consensus on how to proceed with inference
- Broadly speaking, I would classify the proposed approaches into five major categories:
  - Marginal approaches
  - Debiasing
  - Sample splitting/resampling
  - Selective inference
  - Knockoff filter

## Setup

- For all of these methods, we will describe the idea behind how they work and then analyze the same set of simulated data for the sake of comparison
- Simulation setup:



 The hdrm package has a function called genDataABN() to simulate data of this type

## Example data

Our example data set for the next several lectures:

- $n = 100, p = 60, \sigma^2 = 1$
- Six variables with  $\beta_j \neq 0$  (category "A"):
  - Two variables with  $\beta_i = \pm 1$ :
  - Four variables with  $\beta_j = \pm 0.5$ :
- Each of the six variables with  $\beta_j \neq 0$  is correlated ( $\rho = 0.5$ ) with two other variables; i.e., there are 12 "Type B" features
- The remaining 42 variables are pure noise,  $\beta_j=0$  and independent of all other variables ("Type N")

```
genDataABN(n=100, p=60, a=6, b=2, rho=0.5, beta=c(1,-1,0.5,-0.5,0.5,-0.5))
```

### KKT conditions

Recall the KKT conditions for the lasso:

$$\frac{1}{n}\mathbf{x}_{j}'\mathbf{r} = \lambda \operatorname{sign}(\widehat{\beta}_{j}) \qquad \text{for all } \widehat{\beta}_{j} \neq 0$$

$$\frac{1}{n}\left|\mathbf{x}_{j}'\mathbf{r}\right| \leq \lambda \qquad \text{for all } \widehat{\beta}_{j} = 0$$

• Letting  $\mathbf{r}_j = \mathbf{y} - \mathbf{X}_{-j} \hat{\boldsymbol{\beta}}_{-j}$  denote the partial residual with respect to feature j, this implies that

$$\frac{1}{n} \left| \mathbf{x}_{j}' \mathbf{r}_{j} \right| > \lambda \quad \text{for all } \widehat{\beta}_{j} \neq 0$$

$$\frac{1}{n} \left| \mathbf{x}_{j}' \mathbf{r}_{j} \right| \leq \lambda \quad \text{for all } \widehat{\beta}_{j} = 0;$$

similar equations apply for MCP, SCAD, elastic net, etc.

## Selection probabilities

ullet Therefore, the probability that variable j is selected is

$$\mathbb{P}\left(\frac{1}{n}\left|\mathbf{x}_{j}'\mathbf{r}_{j}\right|>\lambda\right)$$

- This suggests that if we are able to characterize the distribution of  $\frac{1}{n}\mathbf{x}_j'\mathbf{r}_j$  under the null, we can estimate the number of false selections in the model
- Indeed, this is easy to do in the case of orthonormal design:

$$\mathbb{E}\left|\hat{\mathcal{S}} \cap \mathcal{N}\right| = 2\left|\mathcal{N}\right| \Phi(-\lambda \sqrt{n}/\sigma),$$

where  $\hat{\mathcal{S}}$  is the set of selected variables and  $\mathcal{N}$  is the set of null variables

#### Estimation

- To use this as an estimate, two unknown quantities must be estimated (this should seem familiar):
  - $\circ~|\mathcal{N}|$  can be replaced by p, using the total number of variables as an upper bound for the null variables
  - $\circ \ \sigma^2$  can be estimated by  $\mathbf{r}^T\mathbf{r}/(n-\left|\hat{\mathcal{S}}\right|)$
- This implies the following estimate for the expected number of false discoveries:

$$\widehat{\mathrm{FD}} = 2p\Phi(-\sqrt{n}\lambda/\hat{\sigma})$$

and this to estimate of the false discovery rate:

$$\widehat{\mathrm{FDR}} = \frac{\widehat{\mathrm{FD}}}{\left|\hat{\mathcal{S}}\right|}$$

## Local false discovery rates

Letting

$$z_j = \frac{\frac{1}{n} \mathbf{x}_j^T \mathbf{r}_j}{\hat{\sigma} \sqrt{n}},$$

we therefore have  $z_j \sim N(0,1)$ 

- We could therefore use this set of z-statistics to estimate feature-specific local false discovery rates as well
- Note that in this approach, we are not restricted to variables in the model;  $z_j$  can be calculated for all p features
- This is all assuming an orthonormal design; what about in the general case?

#### General case

• In the non-orthogonal case,

$$\frac{1}{n}\mathbf{x}_{j}^{T}\mathbf{r}_{j} = \beta_{j}^{*} + \frac{1}{n}\mathbf{x}_{j}^{T}\boldsymbol{\varepsilon} + \frac{1}{n}\mathbf{x}_{j}^{T}\mathbf{X}_{-j}(\boldsymbol{\beta}_{-j}^{*} - \widehat{\boldsymbol{\beta}}_{-j})$$

- Broadly speaking, the general idea here is that:
  - For variables like B, the remainder term is not negligible
  - For variables like N, however, the remainder term is negligible, at least under certain conditions
- For this reason, I named these marginal false discovery rates, as it only establishes FDR control for variables marginally independent of the outcome  $(X_j \perp \!\!\! \perp Y)$ , as opposed to conditional approaches that are concerned with conditional independence:  $X_j \perp \!\!\! \perp Y | \{X_k\}_{k \neq j}$

#### Remarks

Focusing on marginal false discoveries has a few advantages:

- Allows straightforward, efficient estimation of the marginal false discovery rate (mFdr)
- Much more powerful: When two variables are correlated, distinguishing between which of them (or none, or both) is driving changes in Y and which is merely correlated with Y is challenging – even more so in high dimensions
- In many applications, discovering variables like B is not problematic

## Theoretical support

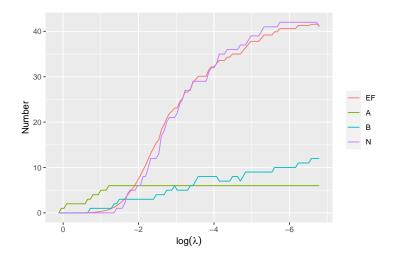
• The design matrix does not have to be strictly orthogonal in order for the proposed estimator to work; let  $\mathcal{A}, \mathcal{N}$  partition  $\{1, 2, \ldots, p\}$  such that  $\beta_j = 0$  for all  $j \in \mathcal{N}$  and the following condition holds:

$$\lim_{n \to \infty} \frac{1}{n} \mathbf{X}' \mathbf{X} = \begin{bmatrix} \Sigma_{\mathcal{A}} & \mathbf{0} \\ \mathbf{0} & \Sigma_{\mathcal{N}} \end{bmatrix}$$

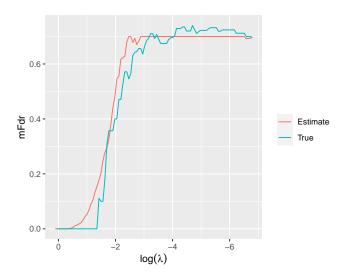
• Theorem: Suppose  $\frac{1}{n}\mathbf{X}_{\mathcal{N}}^T\mathbf{X}_{\mathcal{N}} \to \Sigma_{\mathcal{N}} = \mathbf{I}$ . Then for any  $j \in \mathcal{N}$  and for  $\lambda_n$  such that the sequence  $\sqrt{n}\lambda_n$  is bounded,

$$\frac{1}{\sqrt{n}}\mathbf{x}_j'\mathbf{r}_j \stackrel{\mathrm{d}}{\longrightarrow} N(0,\sigma^2)$$

## mFdr accuracy



# mFdr accuracy (cont'd)

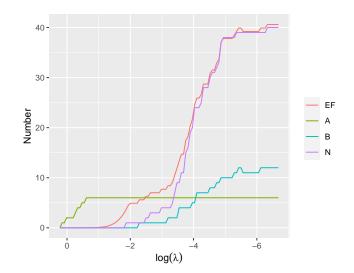


Performance

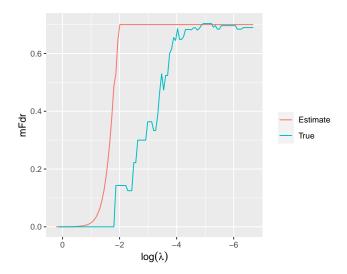
### Correlated noise

- The preceding results are something of a "best case scenario" for the proposed method, since the variables in  ${\cal N}$  were independent
- When the null variables are dependent, the estimator becomes conservative
- The reason for this is that if features are correlated, regression methods such as the lasso will tend to select a single feature and then become less likely to select other correlated features; our calculations do not account for this phenomenon

# mFdr accuracy, highly correlated noise: $\rho_{jk} = 0.8$



## mFdr accuracy, highly correlated noise (cont'd)



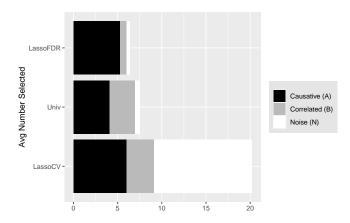
### Comparison

- Being able to estimate mFdr gives us another way of choosing  $\lambda$ : we can choose the smallest value of  $\lambda$  such that  $\mathrm{mFdr}(\lambda) < \alpha$
- For our example data set (uncorrelated noise; nominal FDR = 10%):

	# Selected		
	Α	В	N
Lasso (mFDR)	5	1	0
Univariate	3	2	0
Lasso (CV)	6	3	6

## Comparison (simulation)

A more extensive comparison based on averaging across many simulated data sets:



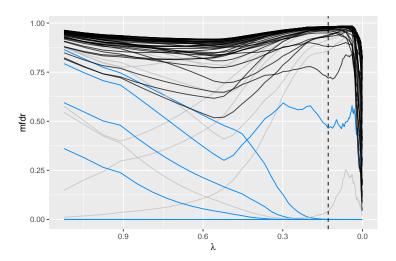
#### Remarks

- Cross-validation gives no control over the number of noise variables selected (and indeed, tends to select a lot of them)
- Univariate approaches give no control over the number of "Type B" variables selected (and also, tend to select a lot of them)
- Using lasso with mFdr control
  - Controls the number of noise variables selected
  - Doesn't necessarily control the number of "Type B" variables selected, but tends not to select many of them (because it's fundamentally a regression-based approach)

### Tension between selection and prediction

- As we saw in our theory lectures, there tends to be a tension between variable selection and prediction, at least for the lasso: values of  $\lambda$  that are optimal for prediction let in too many false positives
- Conversely, if we select  $\lambda$  so as to limit the number of false positives, the resulting model has quite a bit of bias prediction and estimation suffer
- By providing feature-specific inference, local false discovery rates alleviate this tension: we can select the optimal predictive model, but still have a way of quantifying which features are likely to be false discoveries

### Local mfdr



#### summary

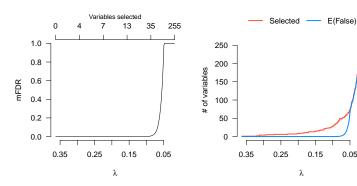
```
> summary(fit, lambda=cvfit$lambda.min)
 Nonzero coefficients : 15
 Expected nonzero coefficients: 7.63
 Average mfdr (15 features) : 0.509
   Estimate z mfdr Selected
A2 -0.84685 -9.717 < 1e-04
A1 0.81777 9.672 < 1e-04
A6 -0.43587 -5.995 < 1e-04
A4 -0.43932 -5.368 < 1e-04
A3 0.34224 4.731 0.00073461
```

## summary (cont'd)

```
. . .
B9
     0.22940 3.776 0.04086459
                                       *
A5
    0.15792 2.870 0.46432327
                                       *
N2
    0.10933 2.462 0.71964287
    -0.06820 -2.088 0.85746822
В3
                                       *
N10
     0.06209 1.992 0.87987162
B10
     0.04748 1.800 0.91334279
N16 -0.03281 -1.664 0.93020634
N41
     0.02980 1.619 0.93487916
                                       *
N6
    -0.02600 -1.567 0.93976257
N34 -0.01123 -1.446 0.94926944
                                       *
```

## Breast cancer data (n = 536, p = 17, 322)

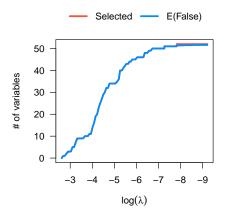
```
mfdr(fit)
plot(mfdr(fit))
```



We can select quite a few variables ( $\approx 50$ ) with a low mFdr

# SOPHIA (n = 292, p = 705, 969)

#### A GWAS example



No features can be selected with any confidence that they are not false inclusions

#### Conclusions

- Marginal false discovery rates are a useful tool for assessing the reliability of variable selection in penalized regression models
- The simplicity of the estimator makes it (a) available at minimal added computational cost and (b) very easy to generalize to new methods
- Some issues to be aware of, though:
  - Only controls FDR in the marginal sense (i.e., not for all  $\beta_j = 0$ )
  - Becomes conservative when noise features are highly correlated
- Local false discovery rates provide a way to select prediction-optimal models without worrying about the number of false selections