

# The Kaplan-Meier Estimator

Patrick Breheny

September 4

# Introduction

- The likelihood construction techniques that we introduced last week can be used to estimate the survival/hazard/distribution functions for any parametric model of survival time
- As I have alluded to in the past, however, the distributions of survival times are often difficult to parameterize
- Our goal for today is to develop *nonparametric* estimates for these distributions, with a particular emphasis on the survival function  $S(t)$

# Empirical survival function

- In the absence of censoring, estimating  $S(t)$  would be straightforward
- We could simply use the empirical survival function

$$\hat{S}(t) = \frac{\#\{i : t_i > t\}}{n}$$

- With censored observations, however, we don't always know whether  $\tilde{T}_i > t$  or not

# Nonparametric likelihood

- As we discussed last week, likelihood provides a natural way to proceed with inference in the presence of censoring
- The likelihood of a survival function  $S$  given observed, right-censored data is

$$\begin{aligned} L(S|\text{Data}) &= \prod_{i=1}^n \mathbb{P}(T_i = t_i)^{d_i} \mathbb{P}(T_i > t_i)^{1-d_i} \\ &= \prod_{i=1}^n \{S(t_i^-) - S(t_i)\}^{d_i} S(t_i)^{1-d_i} \end{aligned}$$

- This expression is a bit different from the likelihoods we saw last week

# Nonparametric likelihood (cont'd)

- In particular, it is not the likelihood of a parameter, but of a generic survival function  $S$
- The set of possible values we must consider is not just an interval of parameter values, but rather the entire set of all possible survival functions
- This is the basic idea of *nonparametric* statistics: rather than specify a parametric form for  $S(\cdot|\theta)$  and carry out inference concerning  $\theta$ , we adopt procedures that deal directly with  $S$  itself

# Estimating $S$

- For today, we will focus on the question of estimating  $S$
- A natural estimate is to choose the value of  $S$  that maximizes  $L(S)$ ; this is the nonparametric maximum likelihood estimator
- In other words, we must determine, out of the set of all possible survival functions, which function maximizes  $L(S)$

# Estimating $S$ : First steps

- This might sound daunting, but it turns out to be easier than you would think
- Let's begin by making two observations that greatly restrict the possible values of  $S$  that we must consider
  - In order to maximize the likelihood,  $S$  must put positive point mass at any time  $t$  at which a subject was observed to fail (otherwise  $S(t_i^-) - S(t_i)$  would be zero)
  - In order to maximize the likelihood,  $S$  cannot put any probability at times other than those at which subjects were observed to fail (redistributing that probability to next failure time would always increase the likelihood)
- Thus, we really only need to determine how much point mass to put at each observed failure time

# Rewriting in terms of observed failure times

- Since the observed failure times are so critical here, let's rewrite the problem in terms of the observed failure times  $0 = t_0 < t_1 < t_2 < \dots < t_J < t_{J+1} = \infty$ , and let

$d_j \equiv \#$  of failures at time  $t_j$

$n_j \equiv \#$  at risk at time  $t_j^-$

$c_j \equiv \#$  censored during the interval  $[t_j, t_{j+1})$

- In terms of this new notation, we can rewrite the earlier likelihood as

$$L(S) = \prod_{j=1}^J \{S(t_j^-) - S(t_j)\}^{d_j} S(t_j)^{c_j}$$



Solving for  $\hat{\lambda}$ 

- Next, let's rewrite the likelihood in terms of the hazard components,  $\hat{\lambda}_1, \dots, \hat{\lambda}_J$
- Doing so yields

$$\begin{aligned} L(\boldsymbol{\lambda}) &= \prod_j \left\{ \lambda_j^{d_j} \prod_{k=1}^{j-1} (1 - \lambda_k)^{d_j} \prod_{k=1}^j (1 - \lambda_k)^{c_j} \right\} \\ &= \prod_j \lambda_j^{d_j} (1 - \lambda_j)^{n_j - d_j} \end{aligned}$$

# Solving for $\hat{\lambda}$ (cont'd)

- Thus, the joint likelihood for  $\lambda$  consists of  $j$  separate components in which  $\lambda_j$  appears only in the  $j$ th component
- Furthermore, each component is equivalent to a binomial likelihood, so

$$\hat{\lambda}_j = d_j/n_j$$

and

$$\begin{aligned}\hat{S}(t) &= \prod_{t_j \leq t} (1 - \hat{\lambda}_j) \\ &= \prod_{t_j \leq t} \frac{n_j - d_j}{n_j}\end{aligned}$$

# Kaplan-Meier estimator

- The estimator on the previous slide was originally proposed by Kaplan and Meier in 1958, and is known as the Kaplan-Meier estimator (or product limit estimator, which is the name Kaplan and Meier proposed)
- This approach has come to be – by far – the most common way of estimating and summarizing survival curves
- The approach is so widespread, in fact, that Kaplan & Meier's original paper is the most highly cited paper in the history of statistics, and the 11th most highly cited paper in all of science

## GVHD study

- To illustrate how Kaplan-Meier estimation works, let's apply it to a study involving graft-versus-host disease (GVHD) in bone marrow transplant recipients
- The patients in the study have a condition called severe aplastic anemia, in which the bone marrow produces an insufficient number of new blood cells
- These patients were given a bone marrow transplant from a compatible family member
- A serious complication of bone marrow transplantation is GVHD, in which the immune cells produced by the new bone marrow recognize the recipient as a foreign body and mount an attack

## GVHD study (cont'd)

- To ward off GVHD, the recipients were randomized to receive one of two drug combinations:
  - Methotrexate (MTX)
  - Methotrexate and cyclosporine (MTX + CSP)
- The goal of the study is to determine whether treatment affected the occurrence of GVHD and if so, which treatment is superior

## Data (by subject)

- Like elsewhere in statistics, survival data is typically organized with each individual subject occupying a row and the outcome and various covariates occupying the columns of the data set
- One difference, however, is that in survival analysis, two columns are required to denote the outcome ( $t_i$  and  $d_i$ ):

Group	Time	Status
MTX+CSP	3	No
MTX+CSP	8	Yes
MTX+CSP	10	Yes
MTX+CSP	12	No
MTX+CSP	16	Yes
	...	

# Data (by time)

As we saw in the derivation of the KM estimator, however, for the purposes of analysis it is often helpful to re-express the data in terms of the observed failure times:

Therapy	Time	GVHD	$t$	$n(t)$	$d(t)$
			0	32	0
MTX+CSP	3	No	3	32	0
MTX+CSP	8	Yes	4	31	0
MTX+CSP	10	Yes	8	31	1
MTX+CSP	12	No	9	30	0
MTX+CSP	16	Yes	10	30	1
	...		16	28	1
				...	

# MTX alone group

In the MTX alone group,

Therapy	Time	GVHD			
MTX	9	Yes			
MTX	11	Yes	<i>t</i>	<i>n(t)</i>	<i>d(t)</i>
MTX	12	Yes	0	32	0
MTX	20	Yes	9	32	1
MTX	20	Yes	11	31	1
MTX	22	Yes	12	30	1
MTX	25	Yes	20	29	2
MTX	25	Yes	22	27	1
MTX	25	No	25	26	2
MTX	28	Yes	28	23	2
MTX	28	Yes	...		
...					



$\hat{S}(t)$ : MTX + CSP

$t$	$n(t)$	$d(t)$	$t$	$\frac{n(t)-d(t)}{n(t)}$	$\hat{S}(t)$
0	32	0	0	1	1
8	31	1	8	30/31	.968
10	30	1	10	29/30	.935
16	28	1	16	27/28	.902
...				...	

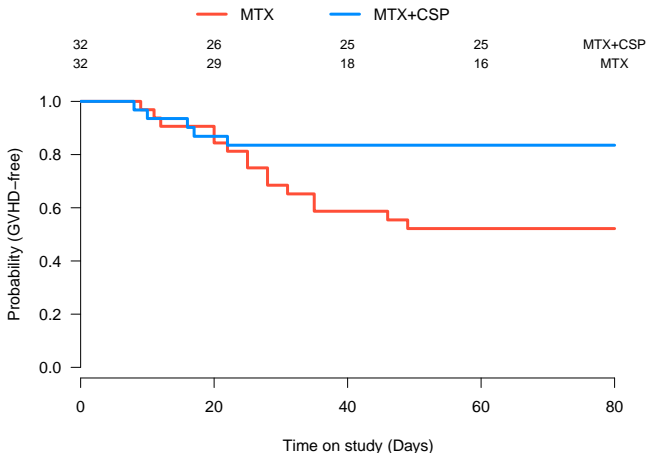
$\hat{S}(t)$ : MTX alone

In the MTX group,

$t$	$n(t)$	$d(t)$	$t$	$\frac{n(t)-d(t)}{n(t)}$	$\hat{S}(t)$
0	32	0	0	1	1
9	32	1	9	31/32	.969
11	31	1	11	30/31	.938
12	30	1	12	29/30	.906
20	29	2	20	27/29	.844
22	27	1	22	26/27	.812
25	26	2	25	24/26	.750
28	23	2	28	21/23	.685
...			...		

# Kaplan-Meier curve: GVHD

The result of all these calculations is usually summarized in a plot called a *Kaplan-Meier curve*:



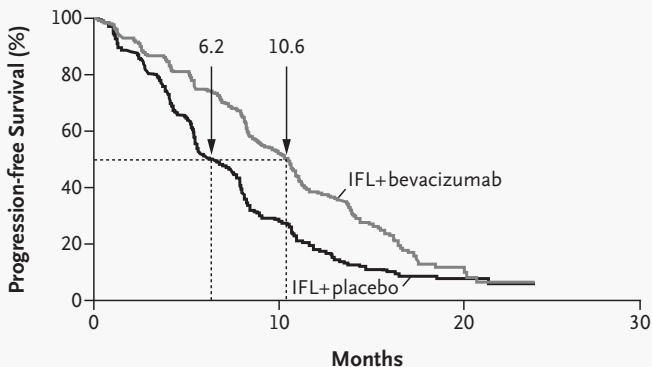
## Summary statistics

- Summary statistics for time-to-event data are typically derived from the Kaplan-Meier estimates
- For example, in this study we might report estimates of the probability of remaining GVHD-free at 60 days of 84% in the MTX+CSP group and 52% in the MTX alone group
- This can be obtained simply by reading the Kaplan-Meier curve “vertically”
- One can also read the Kaplan-Meier curve “horizontally” to obtain estimates of quantiles

# Median survival times

- One quantile of particular interest is the median; i.e., the time at which the survival function drops below 0.5
- In the case where death is the outcome, this is known as the *median survival time* and is almost always reported (if it can be estimated)
- For our GVHD example, the median time to event cannot be estimated since  $\hat{S}(t)$  never reaches 0.5; to see what the idea, though, let's briefly turn to data from a clinical trial of a cancer drug called Avastin

# Kaplan-Meier curve: Avastin study



**No. at Risk**

IFL+bevacizumab	402	269	143	36	6	0
IFL+placebo	411	225	73	17	8	0

# The survival package

- Finally, let's discuss the R functions for constructing the Kaplan-Meier estimate and plotting KM curves
- In this course, we will make extensive use of the survival package in R
- The package is bundled by default with R, meaning that you do not have to install it, although you will have to load it with

```
library(survival)
```

before you can use it

# Surv objects

The survival package has a construct called a Surv object to handle survival outcomes, which are one entity but with two components ( $t_i$  and  $d_i$ ):

```
> S <- with(Data, Surv(Time, Status))
> class(S)
[1] "Surv"
> head(S)
[1] 3+ 8 10 12+ 16 17
> head(S[,1])
[1] 3 8 10 12 16 17
> head(S[,2])
[1] 0 1 1 0 1 1
```



# survfit

- The function in `survival` for constructing Kaplan-Meier estimates is called `survfit`:

```
fit <- survfit(S~Data$Group)
```

where `S` is a `Surv` object

- `S` does not have to be constructed ahead of time; this also works (and is probably better coding practice):

```
fit <- survfit(Surv(Time, Status)~Group, Data)
```

# Summarizing the survfit object

By printing the object, we get a rough summary of each group, although the summary revolves around the median, which in our case cannot be estimated:

```
> fit
```

	n	events	median	0.95LCL	0.95UCL
Group=CSP	32	15	NA	35	NA
Group=CSP+MTX	32	5	NA	NA	NA

Provided they can be estimated, we would see the median survival time in each group, along with upper and lower 95% confidence interval bounds (we'll discuss how those are calculated in the next lecture)

## Summarizing the survfit object (cont'd)

To find out more about the KM estimates at specific times, we can use the `summary` function:

```
> summary(fit, time=40)
      Group=CSP
  time  n.risk  n.event  survival  std.err  95%LCL  95%UCL
40.000  18.000  13.000    0.587    0.088  0.437   0.788

      Group=CSP+MTX
  time  n.risk  n.event  survival  std.err  95%LCL  95%UCL
40.000  25.000  5.0000   0.8353   0.0674  0.7131  0.9784
```

# plot.survfit

- Once the Kaplan-Meier curve has been estimated, it can be plotted in a straightforward manner:

```
plot(fit)
```

- Some useful options to be aware of are
  - `mark.time`: Marks the times at which observations were censored (default: TRUE)
  - `xmax`: Maximum time at which to plot  $\hat{S}(t)$
  - `xscale`: Set this to 365.25 to get curves displayed in years instead of days, and so on

## More attractive-looking plots

- Unfortunately, the default `survival` plots are not particularly nice-looking, and don't provide options for adding the number at risk
- As an alternative, I'm providing:

```
Plot(fit)  
nrisk(fit)
```

which I used to make the plot on slide 19; note that you may need to adjust the margins to avoid the number at risk conflicting with the legend (there are also various R packages available)

- Next time, we'll discuss confidence bands for Kaplan-Meier curves